

MATH 156.3 — TEST III SOLUTIONS

1

The Central Limit Theorem (CLT) says that
 for data from an SRS of size n (problem says SRS and that $n = 500$)
 out of a population where
 the population proportion of those who say yes to some question (here the question is “do you
 perform regular volunteer service?”) is π (problem says $\pi = .25$),
 we have
 the sample proportion \hat{p} is approximately Normal with mean π and standard deviation $\sqrt{\frac{\pi(1-\pi)}{n}}$, *i.e.*,

$$\hat{p} \text{ is } N\left(\pi, \sqrt{\frac{\pi(1-\pi)}{n}}\right) = N(.25, .01936)$$

So $P(\hat{p} > .282) = \text{normalcdf}(.282, 9999, .25, .01936) = .0513$.
 and this is the probability that the sample proportion will be greater than 28.2%.

2

- (a) Let
 Y be the RV “distance to the sun”
 X be the RV “planet number, counting from Mercury = 1”
 for
 the OUs “planets”
 out of
 the population “planets in the Solar System.”
 The LSRL has equation

$$\hat{y} = a + bx \text{ where } b = r \frac{S_Y}{S_X} \text{ and } a = \bar{Y} - b\bar{X}.$$

Here

r is the correlation coefficient of X and Y
 S_X (respectively, S_Y) is the standard deviation of the X data (resp., the Y data)
 \bar{X} (resp., \bar{Y}) is the mean of the X data (resp., the Y data)

To compute these, we put the X data in L1 and the Y data in L2, do **1-Var Stats L1** for S_X and \bar{X} ,
 similarly **1-Var Stats L2** for S_Y and \bar{Y} , and then **LinReg(a+bx) L1,L2** for r . [Actually, that last one gives
 the complete a and b , too.]

This yields $a = -1125.64$ and $b = 445.55$; so

$$\hat{y} = -1125.64 + 445.55x$$

- (b) We want the predicted value of Y – which is \hat{y} – if X is 4.5. Putting $x = 4.5$ into the above eqn of the LSRL
 gives $\hat{y} = -1125.64 + 445.55 \cdot 4.5 = 879.3361$ (\hat{y} in millions of miles).

(c)

The form is not a big random scatter, there is a clear “U”-shape to the residual graph. Therefore the
 linear model is not very good!

- (d) Following the problem instructions we make L3, then we compute the LSRL using the calculator command
 “**LinReg(a+bx) L1,L3**”. This yields $\hat{y} = a + bx = 1.24593 + .27063x$.
 Here is the new residual plot:

There is not much pattern to this residual plot, it goes up and down fairly randomly. It therefore seems that a linear model for *Log of planetary distance vs. planet number* is a fairly good fit.

- (e) The LSRL equation in (d) above gives a predicted *Log of planetary distance* for the asteroid belt (using the planet number 4.5 as in (b)) of: $\hat{y} = a + bx = 1.24593 + .27063 \cdot 4.5 = 2.463765$.
Therefore, as the problem explains, the actual planetary distance of the asteroid belt is predicted by this model to be $10^{2.463765} = 290.914$ million miles.

- 3 (a) There are really two studies here, one for *Sports Illustrated* and one for *Soap Opera Digest*. In both studies, the OUs are “pages of the magazine” and
the RV is binary, categorical, asking “does that page have ads?”
(b) For both, we use the formula

$$\hat{p} \pm z^* \sqrt{\frac{\hat{p}(1 - \hat{p})}{n}}$$

where

- \hat{p} is the sample proportion (our best guess for π),
- z^* is the critical value for the desired confidence level.

Since the problem says to use 95% confidence level, the table gives $z^* = 1.96$.

For *Sports Illustrated*, the pages of the 9/13/1999 issue (which we are told to consider as an SRS of pages out all *SI* issues), numbered 116, out of which 54 had ads. For the 09/14/1999 *Soap Opera Digest* there were 28 pages with ads out of 130 total pages. So the sample proportions are

$$\hat{p}_{si} = 54/116 = .4655 \quad \text{and} \quad \hat{p}_{sod} = 28/130 = .2154.$$

So

the *SI* CI is $.4655 \pm 1.96\sqrt{.4655 \cdot .5345/116} = .4655 \pm .0908$, or (.3747, .5563) [interval notation]
and

the *SOD* CI is $.2154 \pm 1.96\sqrt{.2154 \cdot .7846/130} = .2154 \pm .0707$, or (.1447, .2861).

- (c) The press might say something like
“In *Sports Illustrated*, 46% of the pages have advertisements, with a sampling error of plus or minus 9%. For *Soap Opera Digest*, 21 ± 7% of pages have ads.”
(d) Yes, each interval contains the respective \hat{p} – in fact, the \hat{p} is the center of each interval, which has the form $\hat{p} \pm MoE$
As for the population proportion μ , we have no way of knowing if it will be in any particular CI, although we expect our method is “correct” (*i.e.*, does contain μ) approximately 95% of the time.
(e) “95% confidence” means that
If we repeatedly sample and recalculate new \hat{p} ’s and then new CIs, using the same sample size, confidence level, and CI formulæ each time, then 95% of these new CI will contain the one, fixed (unknown) population proportion μ .

- 4 (a) The formula for CIs has a piece which is the margin of error: $MoE = z^* \sqrt{\frac{\pi(1 - \pi)}{n}}$

so we need to solve $.01 = 2.576\sqrt{\frac{.15 \cdot .85}{n}}$ for n ; here we found

the $z^* = 2.576$ in the table corresponding to a confidence level of 99%

and the problem said to use the 2005/6 figure of 15% as a starting point for the calculations for the following year. But then

$$n = .15 \cdot .85 \left(\frac{2.576}{.01}\right)^2 = 8460.614.$$

Actually, since it is ridiculous for n not to be a whole number, we should say that

for any whole number $n \geq 8461$ the margin of error will be less than 1% with 99% confidence.

- (b) You could either use a
lower confidence level – *e.g.*, with 90% confidence, z^* would become 1.645 and so n would be
 $n \geq .15 \cdot .85 \left(\frac{1.645}{.01}\right)^2 = 3450.182$ (so n any whole number ≥ 3451)

or

a wider target margin of error – *e.g.*, with a goal of $MoE = .02$, we would need only

$$n \geq .15 \cdot .85 \left(\frac{2.576}{.02}\right)^2 = 2115.154 \text{ (so } n \text{ any whole number } \geq 2116).$$

- (c) It is important whenever you calculate a CI that your sample size n and sample proportion \hat{p} always obey

$$n\hat{p} \geq 10 \quad \text{and} \quad n(1 - \hat{p}) \geq 10.$$

So after you change n and resample (thus giving a new \hat{p}), you must recheck these conditions!